Time Series Analysis

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Class 2

Dow Jones

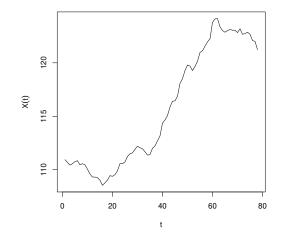


Figure: Dow Jones: average growth.

differenze prime del DJ

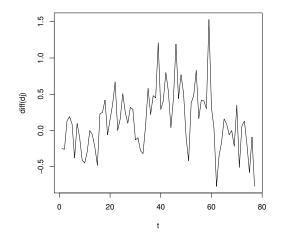


Figure: First differences of Dow Jones: stationary.

- Dow Jones is a stock index that allows to evaluate the overall performance of stock markets.
- It is approximately computed as an average of the 30 most capitalized stocks (although does not account for the different weights of those stocks.



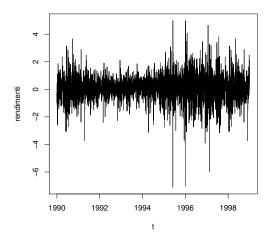


Figure: S&P 500 returns from 1990 to 1999. High volatility.

- Standard and Poor's is a financial company that discloses some indexes about the US financial market.
- The S&P 500 is computed as a weighted arithmetic mean of 500 stocks of US high capitalized companies.
- Returns are computed: $\log \frac{p_t}{p_{t-1}} = \log p_t \log p_{t-1}$.

Quotazioni titolo Italgas

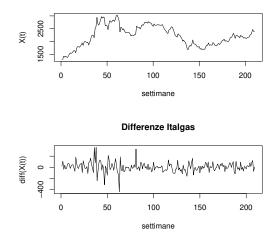


Figure: Close prices for ITALGAS between July 1985 and June 1989. Series and difference series.

- Stationarity represents a link among past, present and future.
- Intuitively, a process {X_t, t = 0, 1...} is said to be stationary if it possesses statistical properties similar to the shifted process {X_{t+h}, t = 0, 1...}, where h is a positive or negative integer.
- Def. A process is said to be strong stationary if the joint distribution $(X_{t_1}, \ldots, X_{t_k})$ is the same of the joint distribution $(X_{t_{1+h}}, \ldots, X_{t_{k+h}})$ for any h, k > 0.
- This property represents a kind of homogeneity on the probabilistic structure of the process with respect to time.
- Example. We have a strong stationary process if $(X_{1985}, X_{1986}, X_{1987})$ has the same distribution as $(X_{1990}, X_{1991}, X_{1992})$, that is $(X_{1985+5}, X_{1986+5}, X_{1987+5})$.

- However, such property is very difficult to be verified !
- We prefer to require a lighter concept of homogeneity based on the second moment.
- A process X_t is said to be weakly stationary if:
- **1** $\mathbb{E}(X_t) = \mu$, the expected value is constant $\forall t$,
- 2 $\mathbb{V}ar(X_t): \gamma(0) = \sigma^2 < \infty$, the variance is constant $\forall t$,
- Ov (X_t, X_{t+h}) = γ(h), the variance at lag h only depends on the distance h, not on t, ∀t.

- An immediate implication is the fact that the correlation structure of the variables does not change in time with respect to the same lag *h*.
- Strong stationary \implies Weak stationary. Viceversa not true.
- The first two moments do not identify a process except for the Gaussian case. Therefore, for a Gaussian process strong and weak stationary property coincide.

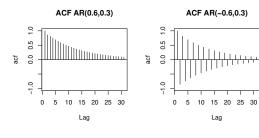
• The autocovariance function of X_t is defined as

$$\gamma(h) = Cov(X_{t+h}, X_t), \qquad h \in \mathbb{R}.$$

• The autocorrelation function (ACF) of X_t is defined as

$$ho(h) = rac{\gamma(h)}{\gamma(0)} = Corr(X_{t+h}, X_t), \qquad h \in \mathbb{R}.$$

- The ACF measures the correlation (linear dependence) between values of the process at different lags *h*, and it indicates the amplitude and length of the memory of the process.
- γ(h) = γ(−h) and ρ(h) = ρ(−h), i.e., the autocovariance function and ACF are even functions (symmetric respect to zero) as the distance in time between X_t e X_{t−h} is the same as X_t e X_{t+h}. For this reason the ACF is usually plotted for positive lags.
- Since $\rho(h)$ is a correlation, it follows $-1 < \rho(h) < 1$.
- Intuition suggests that for a stationary process ρ(h) → 0 (fast or slow) as h increases, otherwise the process would explode.



ACF AR(-0.6,-0.3)

ACF AR(0.6,-0.3)

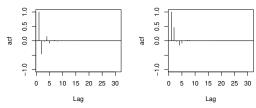


Figure: Examples of ACF.

- Besides the ACF a quantity of interest is the correlation between X_t e X_{t+h} netting for the effect of intermediate variables $X_{t+1}, X_{t+2}, \ldots, X_{t+h-1}$. This is the partial autocorrelation function (PACF).
- The PACF measures the autocorrelation between $X_t \in X_{t+h}$ after removing their linear dependence with the other intermediate variables (recall the partial correlation coefficient in multiple regression).

It takes to compute

$$\phi_{hh} = Corr(X_t, X_{t+h}|X_{t+1}, X_{t+2}, \dots, X_{t+h-1}).$$

 PACF can be derived as follows: consider a regression model where the dependent variable X_{t+h} is regressed against X_{t+h-1}, X_{t+h-2}, ... X_t, i.e.,

$$X_{t+h} = \phi_{h1}X_{t+h-1} + \phi_{h2}X_{t+h-2} + \ldots + \phi_{hh}X_t + e_{t+h},$$

where ϕ_{hj} represents the parameter of the regression of X_{t+h} with respect to the variable X_{t+h-j} , and e_{t+h} is the shock uncorrelated with X_{t+h-j} for $j \ge 1$.

• Note that we are considering a zero mean process.

• Multiplying both sides by X_{t+h-j} and taking the expected values, we get:

$$\gamma(j) = \phi_{h1}\gamma(j-1) + \phi_{h2}\gamma(j-2) + \ldots + \phi_{hh}\gamma(j-h)$$

thus,

$$\rho(j) = \phi_{h1}\rho(j-1) + \phi_{h2}\rho(j-2) + \ldots + \phi_{hh}\rho(j-h).$$

• For *j* = 1, 2, ..., *h* we obtain the following system of equations, known as Yule-Walker equations:

$$\rho(1) = \phi_{h1}\rho(0) + \phi_{h2}\rho(1) + \ldots + \phi_{hh}\rho(h-1)$$

$$\rho(2) = \phi_{h1}\rho(1) + \phi_{h2}\rho(0) + \ldots + \phi_{hh}\rho(h-2)$$
(1)

$$\rho(h) = \phi_{h1}\rho(h-1) + \phi_{h2}\rho(h-2) + \ldots + \phi_{hh}\rho(0)$$

• It can be shown that, after some computations for $h=1,2,\ldots$ we get: $\phi_{11}=
ho(1)$

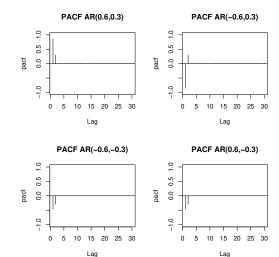
$$\phi_{22} = \frac{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & \rho(2) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{vmatrix}}$$

$$\phi_{33} = \frac{\begin{vmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ \rho(2) & \rho(1) & \rho(3) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ \rho(2) & \rho(1) & 1 \end{vmatrix}}$$

. . .

$$\phi_{hh} = \begin{vmatrix} 1 & \rho(1) & \rho(2) & \dots & \rho(h-2) & \rho(1) \\ \rho(1) & 1 & \rho(1) & \dots & \rho(h-3) & \rho(2) \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ \rho(h-1) & \rho(h-2) & \rho(h-3) & \dots & \rho(1) & \rho(h) \end{vmatrix}$$
$$\begin{vmatrix} 1 & \rho(1) & \rho(2) & \dots & \rho(h-2) & \rho(1) \\ \rho(1) & 1 & \rho(1) & \dots & \rho(h-3) & \rho(2) \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ \rho(h-1) & \rho(h-2) & \rho(h-3) & \dots & \rho(1) & 1 \end{vmatrix}$$

• ACF and PACF are often unknown. We will show how to estimate them. Clearly, to estimate $\hat{\phi}_{hh}$ it takes to estimate $\hat{\rho}(h)$.



Lag

Figure: Examples of PACF.

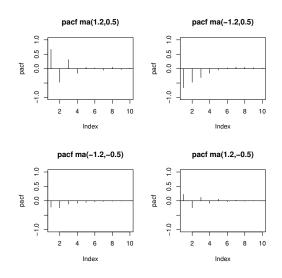


Figure: Examples of PACF.

- Given a stationary process it is possible to compute a unique ACF and PACF.
- We can ask the following: given an ACF is the process that possess that ACF unique ?
- The answer is negative, in general because the ACF does not fully characterize a process. It can be shown that there may exists more processes with same ACF
- The answer is positive, instead, if besides the stationarity we require the invertibility conditions.

- Invertibility relates to the possibility of expressing the process X_t as a function of past random variables.
- More formally, a process {X_t} is said to be invertibile if it exists a linear function h(·) and a white noise {ε_t}, such that for each t:

$$X_t = h(X_{t-1}, X_{t-2}, \cdots) + \epsilon_t.$$

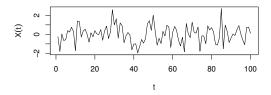
- We will discuss the role of ACF and PACF in finding the model or the underlying process for the observed time series.
- Depending on the specified model we will have different ACF and PACF.

- Example 1. White Noise. The easiest stationary process is the White Noise: $\{X_t\}$ is a sequence of uncorrelated random variables, $Cov(X_t, X_{t+h}) = 0 \ \forall h \neq 0$, zero mean a variance equal to σ^2 .
- Therefore,

$$ho(h) = 1$$
 per $h = 0$
ho(h) = 0 per $h \neq 0.$

- Note that the PACF and ACF coincide since the components are serially uncorrelated.
- We write $X_t \sim WN(0, \sigma^2)$.
- *WN* is a benchmark to assess if the bserved series shows autocorrelation. That is, if a series is autocorelated we can compare it to the ACF of a White Noise.





Series wn

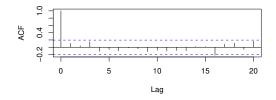


Figure: Representation of a WN series and its ACF.

- Note that Gaussian assumption is not required for a process to be *WN*.
- When adding the Gaussian hypothesis we end up with a sequence of random variables

 $X_t \stackrel{i.i.d}{\sim} N(0, \sigma^2)$